

# Basic calibration functions for analytical chemistry

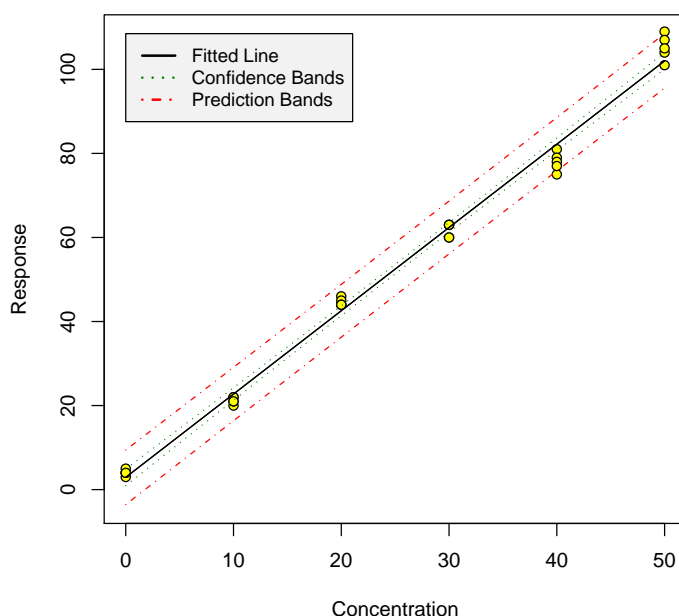
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When calibrating an analytical method, the first task is to generate a suitable model. If we want to use the `chemCal` functions, we will have to restrict ourselves to univariate, possibly weighted <sup>1</sup>, linear regression so far.

Once such a model has been created, the calibration can be graphically shown by using the `calplot` function:

```
> library(chemCal)
> data(massart97ex3)
> attach(massart97ex3)
> m0 <- lm(y ~ x)
> calplot(m0)
```

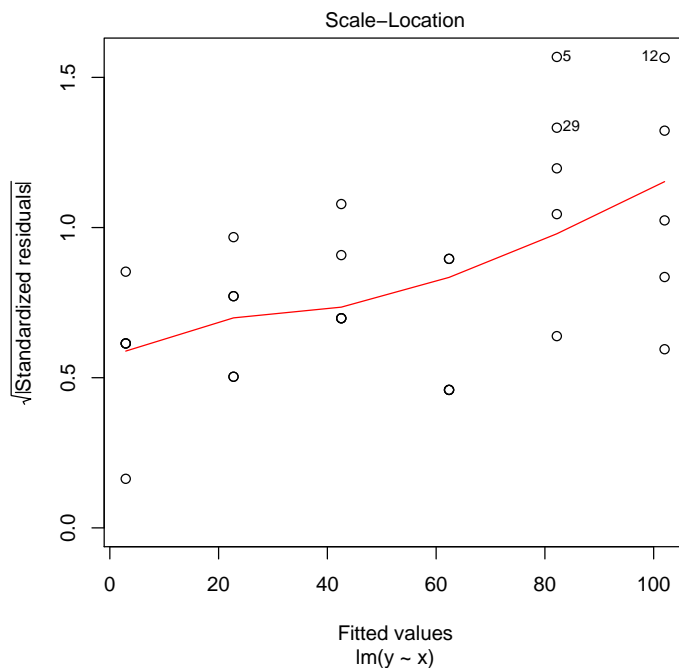


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<sup>1</sup> For the weighted case, the function `predict.lm` would have to be adapted (Bug report PR#8877), in order to allow for weights for the x values used to predict the y values. This affects the functions `calplot` and `lod`.

As we can see, the scatter increases with increasing  $x$ . This is also illustrated by one of the diagnostic plots for linear models provided by R:

```
> plot(m0, which = 3)
```



Therefore, in Example 8 in [1] weighted regression is proposed which can be reproduced by

```
> yx <- split(y, x)
> ybar <- sapply(yx, mean)
> s <- round(sapply(yx, sd), digits = 2)
> w <- round(1/(s^2), digits = 3)
> weights <- w[factor(x)]
> m <- lm(y ~ x, w = weights)
```

Unfortunately, `calplot` does not work on weighted linear models, as noted in the footnote above.

If we now want to predict a new  $x$  value from measured  $y$  values, we use the `inverse.predict` function:

```
> inverse.predict(m, 15, ws = 1.67)
```

```
$Prediction
[1] 5.865367
```

```
$`Standard Error`
[1] 0.892611
```

```

$Confidence
[1] 2.478285

$`Confidence Limits`
[1] 3.387082 8.343652

> inverse.predict(m, 90, ws = 0.145)

$Prediction
[1] 44.06025

$`Standard Error`
[1] 2.829162

$Confidence
[1] 7.855012

$`Confidence Limits`
[1] 36.20523 51.91526

```

The weight **ws** assigned to the measured *y* value has to be given by the user in the case of weighted regression. By default, the mean of the weights used in the linear regression is used.

## Theory for `inverse.predict`

Equation 8.28 in [1] gives a general equation for predicting the standard error  $s_{\hat{x}_s}$  for an *x* value predicted from measurements of *y* according to the linear calibration function  $y = b_0 + b_1 \cdot x$ :

$$s_{\hat{x}_s} = \frac{s_e}{b_1} \sqrt{\frac{1}{w_s m} + \frac{1}{\sum w_i} + \frac{(\bar{y}_s - \bar{y}_w)^2 \sum w_i}{b_1^2 \left( \sum w_i \sum w_i x_i^2 - (\sum w_i x_i)^2 \right)}} \quad (1)$$

with

$$s_e = \sqrt{\frac{\sum w_i (y_i - \hat{y}_i)^2}{n - 2}} \quad (2)$$

where  $w_i$  is the weight for calibration standard *i*,  $y_i$  is the mean *y* value (!) observed for standard *i*,  $\hat{y}_i$  is the estimated value for standard *i*, *n* is the number calibration standards,  $w_s$  is the weight attributed to the sample *s*, *m* is the number of replicate measurements of sample *s*,  $\bar{y}_s$  is the mean response for the sample,  $\bar{y}_w = \frac{\sum w_i y_i}{\sum w_i}$  is the weighted mean of responses  $y_i$ , and  $x_i$  is the given *x* value for standard *i*.

The weight  $w_s$  for the sample should be estimated or calculated in accordance to the weights used in the linear regression.

I adjusted the above equation in order to be able to take a different precisions in standards and samples into account. In analogy to Equation 8.26 from [1] we get

$$s_{\hat{x}_s} = \frac{1}{b_1} \sqrt{\frac{s_s^2}{w_s m} + s_e^2 \left( \frac{1}{\sum w_i} + \frac{(\bar{y}_s - \bar{y}_w)^2 \sum w_i}{b_1^2 \left( \sum w_i \sum w_i x_i^2 - (\sum w_i x_i)^2 \right)} \right)} \quad (3)$$

## References

- [1] Massart, L.M, Vandeginste, B.G.M., Buydens, L.M.C., De Jong, S., Lewi, P.J., Smeyers-Verbeke, J. Handbook of Chemometrics and Qualimetrics: Part A, Elsevier, Amsterdam, 1997